

Sex Differences and Allee Effects Shape the Dynamics of Sex-Structured Invasions

Allison K. Shaw^{*a}, Hanna Kokko^b and Michael G. Neubert^c

^aDepartment of Ecology, Evolution, and Behavior, University of Minnesota, MN 55108 USA

^bDepartment of Evolutionary Biology and Environmental Studies, University of Zurich, Winterthurerstrasse 190, CH-8057 Zurich, Switzerland

^cBiology Department, MS 34, Woods Hole Oceanographic Institution, Woods Hole, MA 02543 USA

^{*}Corresponding author: ashaw@umn.edu

Running headline: Sex-structured invasion dynamics

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Summary

1. The rate at which a population grows and spreads can depend on individual behaviour and interactions with others. In many species with two sexes, males and females differ in key life history traits (e.g. growth, survival, dispersal), which can scale up to affect population rates of growth and spread. In sexually reproducing species, the mechanics of locating mates and reproducing successfully introduce further complications for predicting the invasion speed (spread rate), as both can change nonlinearly with density.
2. Most models of population spread are based on one sex, or include limited aspects of sex differences. Here we ask whether and how the dynamics of finding mates interact with sex-specific life history traits to influence the rate of population spread.
3. We present a hybrid approach for modelling invasions of populations with two sexes that links individual-level mating behaviour (in an individual-based model) to population-level dynamics (in an integrodifference equation model).
4. We find that limiting the amount of time during which individuals can search for mates causes a demographic Allee effect which can slow, delay or even prevent an invasion. Furthermore, any sex-based asymmetries in life history or behaviour (skewed sex ratio, sex-biased dispersal, sex-specific mating behaviours) amplify these effects. In contrast, allowing individuals to mate more than once ameliorates these effects, enabling polygynandrous populations to invade under conditions where monogamously mating populations

would fail to establish.

5. We show that details of individuals' mating behaviour can impact the rate of population spread. Based on our results, we propose a stricter definition of a mate-finding Allee effect, which is not met by the commonly used minimum mating function. Our modelling approach, which links individual- and population-level dynamics in a single model, may be useful for exploring other aspects of individual behaviour that are thought to impact the rate of population spread.

Keywords

integrodifference equations, invasion speed, mate-finding Allee effect, mating functions, sex-biased dispersal, spatial spread

Introduction

1 The key characteristics of a potential invasion are whether a population will be
2 able to establish and, if it does, how fast it will subsequently grow and spread.
3 Often there is a threshold size or density below which a population will be unable
4 to successfully establish ([Volterra, 1938](#); [Tobin *et al.*, 2007](#)). Early theoretical work
5 predicted that, when populations do establish, they should spread at a constant
6 rate, assuming identical individuals move diffusely in a homogeneous environment
7 ([Skellam, 1951](#)). More recent theory has explored how invasion speed can be in-
8 fluenced by a variety of factors, including long-distance dispersal ([Kot, Lewis &](#)
9 [van den Driessche, 1996](#)), population structure ([Neubert & Caswell, 2000](#)), demo-

10 graphic stochasticity (Kot *et al.*, 2004), spatial heterogeneity (Li, Fagan & Meyer,
11 2015), environmental stochasticity (Lewis & Pacala, 2000), temporal variability
12 (Neubert, Kot & Lewis, 2000; Caswell, Neubert & Hunter, 2011; Schreiber &
13 Ryan, 2011), and evolution (Kanarek & Webb, 2010; Phillips, Brown & Shine,
14 2010; Perkins *et al.*, 2013). However, the majority of these invasion models as-
15 sume an asexual population and may fail to capture dynamics of spatial spread in
16 sexually reproducing species for two reasons.

17 First, individuals in sexually reproducing species must locate at least one mate
18 in order to reproduce successfully. These species may experience a mate-finding
19 Allee effect where an individual’s probability of finding mates decreases with lower
20 population density (Dennis, 1989; Wells *et al.*, 1998). This so-called “component
21 Allee effect” (which impacts one ‘component’ of an individual’s fitness) results in
22 a demographic Allee effect if, over some range of density, the overall population
23 growth rate increases with a (small) increase in population density (Stephens,
24 Sutherland & Freckleton, 1999). Since Allee effects tend to have the strongest
25 impact at low population sizes and densities, they are especially likely to affect the
26 viability of endangered species, whose numbers have been reduced (Courchamp,
27 Berec & Gascoigne, 2008), and the ability of invasive species to establish and
28 spread from an initially low density at newly established sites (Keitt, Lewis &
29 Holt, 2001; Taylor & Hastings, 2005).

30 Second, in many cases, males and females differ in key life history traits. Sex
31 differences can occur in age of sexual maturity, reproductive period length and
32 total lifespan (Short & Balaban, 1994; Fairbairn, 2013) Males and females can
33 also differ in their tendency to disperse, and total dispersal distance travelled
34 (Greenwood, 1980; Clarke, Sæther & Røskft, 1997; Miller *et al.*, 2011). Since

the speed of an invasion depends on both demographic and dispersal parameters (Skellam, 1951; Neubert & Caswell, 2000), sex-specific differences in any of these parameters can potentially alter the rate of spread. While the consequences of Allee effects for spatial spread has been studied (Fife & McLeod, 1977; Lewis & Van Den Driessche, 1993; Kot, Lewis & van den Driessche, 1996; South & Kenward, 2001; Taylor & Hastings, 2005; Jerde, Bampfylde & Lewis, 2009), this is rarely done while simultaneously taking into account sex differences in dispersal.

To overcome these limitations, we take advantage of two approaches already developed, that allow us to study spatial spread in sex-structured populations. Miller *et al.* (2011) developed an integrodifference equation model that tracks male and female density separately, allowing the exploration of how sex differences in dispersal and the nature of the mating system jointly influence the spread rate. Although this is a (relatively) mathematically tractable approach, it ignores nonlinearities in mating (introduced by mate-finding difficulties) at low density. Shaw & Kokko (2014, 2015) developed an individual-based simulation model with male and female individuals, allowing them to include a mechanistic mate-finding Allee effect by explicitly modelling the mate finding process. Although this approach allows for inclusion of more biological detail than the approach of Miller *et al.* (2011), it is less mathematically tractable and requires extensive simulation time. These two approaches have each demonstrated that sex-specific dynamics (sex-biased dispersal, mate finding) can dramatically influence a population's spread rate.

Here we present a hybrid approach that couples individual-based simulation and integrodifference equation modelling methods. We model mating dynamics using individual-based simulations from Shaw & Kokko (2014) and then insert

60 these into the integrodifference equation framework from [Miller *et al.* \(2011\)](#) to
 61 describe the growth and spatial spread of a population. We compare our results
 62 using the simulated mating dynamics to results using two different mathematical
 63 approximations of the mating dynamics. We show that sex-specific mating be-
 64 haviour and sex differences in life history each affect the population growth rate,
 65 probability of establishment, and dynamics of spread, and furthermore that mat-
 66 ing and life history effects are interactive. We expect that this hybrid framework
 67 will be a useful tool for future studies that aim to explore how other aspects of
 68 individual behaviour influence population outcomes, particularly in a spreading
 69 context.

70 **Materials and methods**

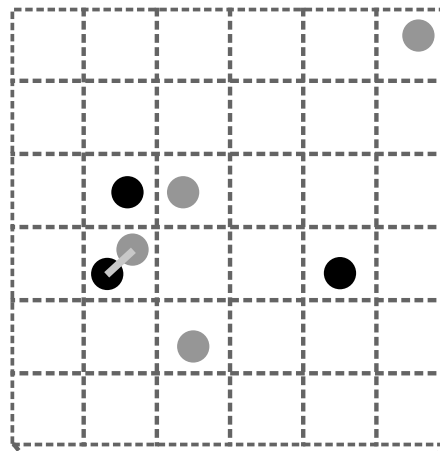
71 Full model code is available from Dryad ([Shaw, Kokko & Neubert, 2016](#)).

72 **Mating dynamics**

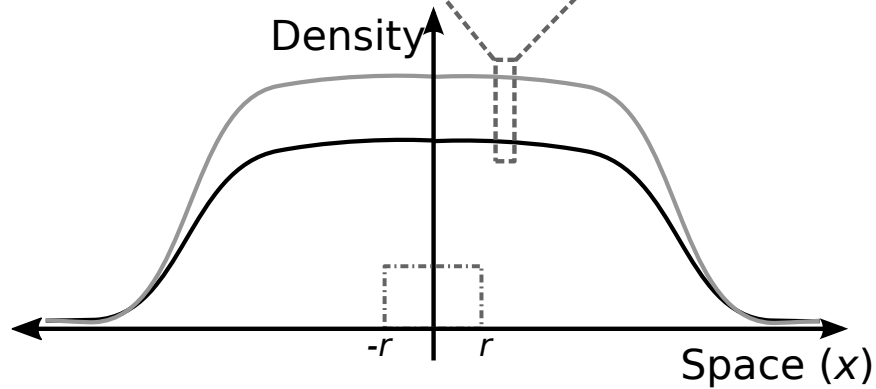
73 We start by describing the mating dynamics that occur locally among f females
 74 and m males (Fig. 1). Let $\mathfrak{B}(f, m)$ be the total number of matings that occur at a
 75 location, i.e. the mating function. We first consider a mating function derived ex-
 76 plicitly from individual-based simulations of the mate finding and mating process.
 77 Then we consider two functions that are approximations of the mating process:
 78 the mass action function and the Hölder mating function. As we show below, as
 79 special cases these also produce simpler formulations such as the minimum mating
 80 function ([Bessa-Gomes, Legendre & Clobert, 2010](#)) where the number of matings
 81 simply equals the number of individuals of the rarer sex.

Figure 1: Schematic showing the two scales of our hybrid model. Mating occurs between f females (black) and m males (grey) at a local scale. A male and female in the same patch can pair (indicated by a line) and mate. (Note that each patch may contain more than one male and/or female.) The density of males and females is tracked at the population scale (here shown as the density of individuals immediately after births occur). The dashed-dotted line shows the initial population density distribution for each simulation.

Local scale
(mating dynamics)



Population scale
(growth, dispersal)



Individual-based functions

The first mating function is based on the individual-based simulation developed by Shaw & Kokko (2014, 2015). We determine the number of monogamous mating pairs that form from f females and m males as follows (we address the case of non-monogamy later). We first randomly distributed f females and m males in patches across a 10-by-10 square lattice with wrapping boundaries. Within each time step ($t = 1, \dots, \tau$) that follows, each individual can make one mate search decision, then form pairs (if possible). Once individuals are part of a pair, they do not search in subsequent time steps. This searching-pairing process is repeated for τ time steps, after which the total number of formed pairs is recorded. Note that each patch can contain any number of males and females (up to m and f , respectively).

In this part of the model, each individual's movement is governed by a mate search function. During the mate search decision, each individual chooses to stay in their patch with a probability that we model using a logistic function of the number of females and males present,

$$p_f = \frac{1}{1 + 0.5e^{-a(R_{FS}-f_i)} + 0.5e^{a(R_{FO}-m_i)}} \quad (1a)$$

$$p_m = \frac{1}{1 + 0.5e^{a(R_{MO}-f_i)} + 0.5e^{-a(R_{MS}-m_i)}} \quad (1b)$$

for females and males respectively. Individuals choose to leave their patch with probability $1 - p_f$ and $1 - p_m$. Here, f_i and m_i are the number of unpaired females and males in the i th patch, R_{FS} and R_{FO} are fixed strategy parameters for a

female's response to individuals of the same and opposite sex, R_{MS} and R_{MO} are fixed strategy parameters for a male's response to individuals of the same and opposite sex, and a is the shape parameter of the strategy as described in [Shaw & Kokko \(2014\)](#). For $a > 0$ (the only case we consider) a female is more likely to stay in a patch that has few females and many males, and a male is more likely to stay in a patch that has many females and few males.

To minimize confusion later on, we refer to movement that occurs during this mating process as 'searching' to distinguish it from the 'dispersal' movement that we include in the full population model below. Here, we consider three scenarios, distinguished by the values of the search strategy parameters: unbiased searching (both males and females search for mates; $R_{\text{FS}} = R_{\text{FO}} = R_{\text{MS}} = R_{\text{MO}} = 1$), male-biased searching (only males search and females are stationary; $R_{\text{FS}} \rightarrow \infty$, $R_{\text{FO}} \rightarrow -\infty$, $R_{\text{MS}} = R_{\text{MO}} = 1$), and female-biased (only females search and males are stationary; $R_{\text{FS}} = R_{\text{FO}} = 1$, $R_{\text{MS}} \rightarrow -\infty$, $R_{\text{MO}} \rightarrow \infty$). The parameter values $R_{\text{JS}} = R_{\text{JO}} = 1$ indicate that an individual of sex j has a 50% probability of leaving a patch with one member of the same sex and one member of the opposite sex, will be more likely to leave patches with more members of the same sex, and less likely to leave patches with more members of the opposite sex. (See [Shaw & Kokko \(2014, 2015\)](#) for graphical examples of these probability functions.)

Individuals search serially, in random order within each time step t . If an individual chooses to leave its patch, it moves randomly to one of the four patches adjacent to its current patch (with wrapping boundaries across the lattice). After all individuals have made a mate search decision, pairs form. In each patch i that contains both females and males, $\min(f_i, m_i)$ pairs are formed. Note that once individuals are paired, they are ignored by non-paired individuals; mate search

126 decisions are only made based on the densities of unpaired males and females.

127 For comparison, we also ran simulations allowing for multiple matings (polygy-
128 nandry). In these simulations, both males and females searched for mates ($R_{\text{FS}} =$
129 $R_{\text{FO}} = R_{\text{MS}} = R_{\text{MO}} = 1$). Here, all mating pairs are temporary, and new pairs can
130 potentially form at every time step. All individuals (whether or not they mated in
131 the past) make a mate search decision at every time step, according to equation
132 (1) which (for this set of simulations) depended on the total number of males and
133 females in the current patch (regardless of mating history). At the end of these
134 simulations, instead of calculating the number of pairs that formed, we counted
135 the number of females who had mated at least once.

136 We ran a total of 1,764,000 simulations. For each of the four scenarios (unbi-
137 ased, male-biased, female-biased, polygynandrous matings), we ran simulations
138 with all factorial combinations of $f = 1, 11, 21, \dots, 191, 200$ females and $m =$
139 $1, 11, 21, \dots, 191, 200$ males for $\tau = 1, 2, 3, \dots, 500$ time steps, to generate four cor-
140 responding mating functions (\mathfrak{B}_{u} , \mathfrak{B}_{m} , \mathfrak{B}_{f} and \mathfrak{B}_{p}) as a function of the number
141 of females (f), number of males (m), and mate search time (τ). We ran 1,000
142 replicates for each combination of parameters and then averaged across replicates.

143 **Mass action function**

144 Next, we consider a second mating function that is an approximation of the pair
145 formation process. This approach is similar to the one used by [Veit & Lewis \(1996\)](#).
146 The change in the number of females, males and pairs as a function of search time

147 can be described by the set of equations

$$\frac{df}{d\hat{\tau}} = -M(f, m) \quad (2a)$$

$$\frac{dm}{d\hat{\tau}} = -M(f, m) \quad (2b)$$

$$\frac{dp}{d\hat{\tau}} = M(f, m) \quad (2c)$$

148 where $M(f, m)$ is the rate of pair formation and $0 \leq \hat{\tau} \leq \tau$. By assuming random
149 searching of males and females, we can model the pair formation process with the
150 law of mass action

$$M(f, m) = cf m = c(f_0 - p)(m_0 - p) \quad (3)$$

151 where f_0 and m_0 are the initial number of males and females present. After a fixed
152 length of time τ , the number of pairs that formed is equal to

$$\mathfrak{B}_a = \begin{cases} \frac{f_0 m_0 \{1 - \exp[(m_0 - f_0)c\tau]\}}{f_0 - m_0 \exp[(m_0 - f_0)c\tau]} & \text{if } f_0 \neq m_0 \\ \frac{n_0^2}{n_0 + (1/c\tau)} & \text{if } f_0 = m_0 = n_0 \end{cases} \quad (4)$$

153 As $\tau \rightarrow \infty$, all possible pairs form and $\mathfrak{B}_a(\tau) = \min(f_0, m_0)$, i.e., the mini-
154 mum mating function. Similar to the individual-based mating function above,
155 we calculated the number of matings for each $f = 1, 11, 21, \dots, 191, 200$ females
156 and $m = 1, 11, 21, \dots, 191, 200$ males for $\tau = 1, 2, 3, \dots, 500$ time steps with $c = 0.01$
157 ($c\tau = 0.01, 0.02, 0.03, \dots, 5$). The end result is the mating function \mathfrak{B}_a as a function
158 of the number of females (f), number of males (m), and $c\tau$.

159 Hölder mean function

160 Finally, we consider an approximation for the number of matings without the pair
161 formation dynamics. Here, the mating function is given by

$$\mathfrak{B}_h(f_t, m_t) = [\beta f_t^\alpha + (1 - \beta)m_t^\alpha]^{1/\alpha} \quad (5)$$

162 with shape parameter $\alpha < 0$ and weight parameter $0 \leq \beta \leq 1$, which describes
163 the relative weighting of females and males. This is a family of weighted means
164 (Hölder means), first suggested to describe mating dynamics by [Hadeler \(1989\)](#),
165 and commonly used as a mating function (including by [Miller *et al.* 2011](#)). For
166 different values of the shape parameter α and the weight parameter β this func-
167 tion encompasses a number of commonly used mating functions. For $\alpha = -\infty$
168 and $\beta = 0.5$ (the minimum mating function) we recover the mass action mat-
169 ing function (eqn. 4) with $\tau \rightarrow \infty$. Similar to the mating functions above, we
170 calculated the number of matings for each $f = 1, 11, 21, \dots, 191, 200$ females and
171 $m = 1, 11, 21, \dots, 191, 200$ males for $\alpha = 0, -0.5, -1, -10, -\infty$ and $\beta = 0.5$. The
172 end result is the mating function \mathfrak{B}_h as a function of the number of females (f),
173 number of males (m), and shape parameter α .

174 Spatial spread framework

175 Next, to describe the growth and spatial spread of the population, we use the
176 two-sex integrodifference equation model developed by [Miller *et al.* \(2011\)](#). We
177 now track the densities of females and males at location x and time t , given by
178 $f_t(x)$ and $m_t(x)$, respectively. Time is discrete (t is an integer) and space is
179 continuous (x is any real number). Although most organisms move in at least

two spatial dimensions of habitat (and indeed we simulate mating dynamics in two dimensions as described above), here we simulate spatial spread in a single dimension for simplicity.

Four processes (dispersal, mating, reproduction, and survival) occur sequentially to determine the number of individuals at the next time step ($t + 1$). First, individuals disperse, with potentially different dispersal kernels for females and males, given by $k_f(x, y)$ and $k_m(x, y)$, respectively. The dispersal kernel describes the probability of an individual from any given starting location y dispersing to any given ending location x . For the purposes of this paper we assume the dispersal kernels only depend on the absolute distance travelled, so we may write the kernel as functions of a single variable, i.e., $k_f(|x - y|)$ and $k_m(|x - y|)$. Second, males and females mate, and the total number of matings at each location x is given by the mating function $\mathfrak{B}(f(x), m(x))$ (described in the previous section). Note that mating occurs locally (Fig. 1). Third, each mating produces ϕ female and μ male offspring. Finally, only a fraction of offspring survive, given by the density-dependent function g . Generally we assume that g decreases with increasing adult population size, so that

$$g(0, 0) = 1 \tag{6a}$$

$$\frac{\partial g}{\partial f} \leq 0 \tag{6b}$$

$$\frac{\partial g}{\partial m} \leq 0 \tag{6c}$$

for all $f, m \geq 0$. These assumptions preclude Allee effects acting through this component of the model. In our simulations, we use the same density dependence

199 function used by [Miller *et al.* \(2011\)](#),

$$g = \frac{b}{b + m + f} , \quad (7)$$

200 where b is the population density at which survival is reduced by 50% due to density
201 dependence. Adults die after reproduction, so generations are non-overlapping.

202 Bringing together all of these processes gives the full model

$$f_{t+1}(x) = \phi g \left[k_f(x) * f_t(x), k_m(x) * m_t(x) \right] \mathfrak{B} \left[k_f(x) * f_t(x), k_m(x) * m_t(x) \right] \quad (8a)$$

$$m_{t+1}(x) = \mu g \left[k_f(x) * f_t(x), k_m(x) * m_t(x) \right] \mathfrak{B} \left[k_f(x) * f_t(x), k_m(x) * m_t(x) \right] \quad (8b)$$

where “ $*$ ” denotes a convolution operator, defined as

$$k(x) * n(x) = \int_{-\infty}^{\infty} k(x - y) n(y) dy . \quad (9)$$

203 We numerically simulated model (8) with the different mating functions (\mathfrak{B}_u ,
204 \mathfrak{B}_m , \mathfrak{B}_f , \mathfrak{B}_p , \mathfrak{B}_a , \mathfrak{B}_h) and examined the invasion dynamics. To run a spatial
205 simulation, we started with an initial density of 0.5 for each sex for $|x| < r$ and 0
206 elsewhere (dashed-dotted line in Fig. 1). Here we used Gaussian dispersal kernels
207 with mean 0 and variance v_f for females and v_m for males, and ran simulations for
208 100 generations. We defined the edge of the population as the first location where
209 the population density was above the threshold value of 0.01. We then calculated
210 the rate of population spread as the difference between the location of the edge of

the population from one time step to the next.

For comparison, we also simulated a simplified version of model (8) without dispersal or space, given by

$$f_{t+1} = \phi g[f_t, m_t] \mathfrak{B}[f_t, m_t] \quad (10a)$$

$$m_{t+1} = \mu g[f_t, m_t] \mathfrak{B}[f_t, m_t] . \quad (10b)$$

For these simulations, we calculated the per capita growth rate as

$$R = \frac{f_{t+1} + m_{t+1} - (f_t + m_t)}{f_t + m_t} \quad (11)$$

in order to examine how the different mating functions affect population growth, varying both the amounts of time allowed for mate searching (τ) in the mating functions, as well as the initial population radius (r).

Hybrid framework

Finally, we link the mating functions developed in the *Mating dynamics* section (described in terms of discrete numbers of individuals and matings) with the spatial spread framework developed in the *Spatial spread framework* section (described in terms of continuous densities), into a hybrid framework. To do so, we first divide each of the mating functions (which give the number of matings as a function of the numbers of males and females) by the area considered (100 patches of 1 unit area each) to get the density of matings as a function of the densities of

males and females. Next, we use linear interpolation to generate estimates of mating density for male and female densities that fall between the values that we simulated. Finally, we generate estimates of mating density for male and female densities that fall outside the range of values that we simulated (below 0.01 or above 2.0). Here we consider two different case: plateauing and extrapolating.

As a default, we assume that the number of matings plateaus at both low and high densities. When either the value of f or m falls below the lowest value in the mating function (0.01) at a location y , we assume that no matings occur at that location (i.e., $\mathfrak{B}(f(y), m(y)) = 0$). When both the values of f and m fall above 2 (the highest value in the mating function) at a location y , we assume that the maximum number of matings found in the mating function occur ($\mathfrak{B}(f(y), m(y)) = \max_{f,m}(\mathfrak{B})$). When just $f > 2$, we determine matings based on the m value and $f = 2$. Similarly, when just $m > 2$, we determine matings based on the f value and $m = 2$.

For comparison, we also considered the case where there is no bound on the number of matings at either low or high density. When either the value of f or m falls below the lowest value in the mating function, we use linear extrapolation with $\mathfrak{B}(0, m) = \mathfrak{B}(f, 0) = 0$ to estimate the matings that occur. When either the values of f or m fall above the highest value in the mating function, we assume that $\min(f, m)$ matings occur.

Results

We first calculated the number of matings according to each mating function, then we examined how each mating function affects the population growth rate in

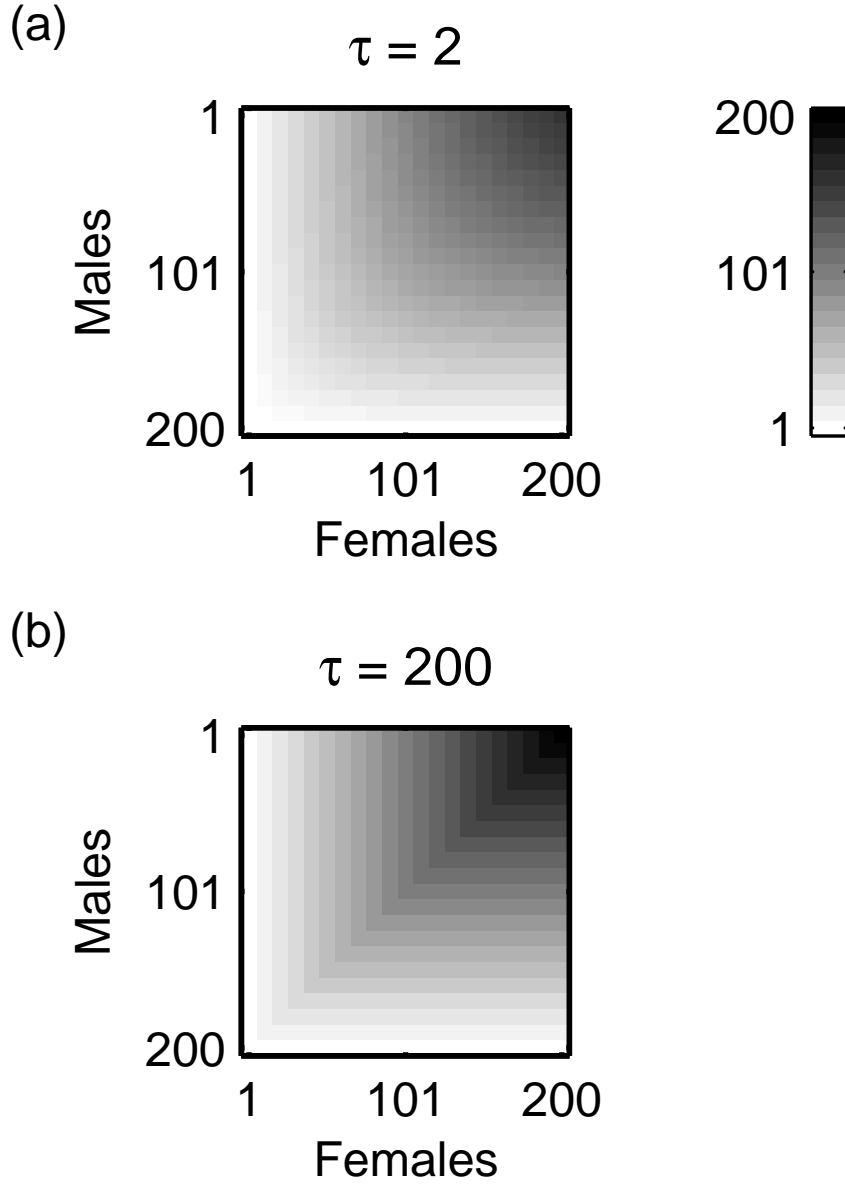
249 a non-spatial model, and finally we determined how each mating function affects
250 the establishment probability and the subsequent rate of spread in the full spatial
251 model.

252 Mating dynamics

253 Intuitively, increasing the number of males, the number of females, or the time
254 allowed for pair formation increased the number of mated pairs that formed. This
255 was true for mating function \mathfrak{B}_u (individual-based mating function with equal
256 searching by males and females), as well as the sex-biased mate searching mating
257 functions, \mathfrak{B}_f and \mathfrak{B}_m (Figs 2, S1a-c). There was also essentially no difference
258 between the number of matings produced by \mathfrak{B}_f and \mathfrak{B}_m (Fig. S1b-c). Sex-biased
259 searching produced the same number of matings as unbiased searching when the
260 sex ratio was skewed, and fewer matings when the sex ratio was close to even
261 (Fig. S1d-e). The mating function \mathfrak{B}_p (polygynandrous mating) approaches f for
262 large τ (Fig. S2a). The mating function \mathfrak{B}_a (mass action) produced essentially
263 the same number of matings as \mathfrak{B}_u (Fig. S2). Although the \mathfrak{B}_h (Hölder means)
264 mating function does not include search time explicitly as a parameter, shifting α
265 more negative produces qualitatively similar results as increasing the search time
266 τ in the individual-based mating function (Figs S1a, S2c). The Hölder means,
267 however, overestimated the number of mates at low densities and for short search
268 times (Fig. S2e).

269 Each mating function meets the criteria described by Courchamp, Berec &
270 Gascoigne (2008) for a mate-finding Allee effect. That is, the female mating rate
271 $P(f, m)$ (or fraction of females that are mated) satisfies the following four proper-

Figure 2: A contour plot of the number of pairs formed as a function of the number of females and males present, for the unbiased mating function (\mathfrak{B}_u), for $\tau = 2$ and $\tau = 200$ (the number of time steps allowed for pair formation). Mating functions were calculated using the mate search shape parameter $a = 5$.



272 ties:

- 273 1. "There is no mating if there are no males: $P(f, 0) = 0$ for any f .
- 274 2. For a given number of females, a female's probability of mating (or mating
275 rate) cannot decrease if the number of males increases: $P(f, m)$ is a non-
276 decreasing function of m for any fixed f .
- 277 3. For a given number of males, a female's probability of mating (or mating
278 rate) cannot increase if the number of females increases: $P(f, m)$ is a non-
279 increasing function of f for any fixed m .
- 280 4. If males greatly outnumber females, mating is virtually certain for females:
281 $P(f, m)$ approaches 1 for a sufficiently large m/f ratio."

282 Note, however, that most verbal definitions of a mate-finding Allee effects describe
283 it as a difficulty or failure to find mates at low density, or a limited amount of
284 time available for mating (Dennis, 1989; Stephens, Sutherland & Freckleton, 1999;
285 Calabrese & Fagan, 2004). Although the criteria above describe how a female's
286 mating probability should change as the density of each males and females change
287 separately, they do not provide a description of how mating probability changes
288 with overall changes in population density (with a fixed sex ratio). Therefore we
289 recommend that a fifth criterion be added:

- 290 5. At low density, equally increasing the number of males and females present
291 should increase a female's probability of mating (or mating rate): $P(af, am) >$
292 $P(f, m)$ for small f and m , and $a > 1$.

293 With this added criterion, the minimum mating function (and the general Hölder
294 means mating function) does not display a mate-finding Allee effect, since a fe-

male's probability of mating remains constant as male and female density increase proportionally. Our classification of the minimum mating function is in line with verbal statements of mate-finding Allee effects, which typically require some form of limitation (due to low population density or short mating time) not present in the minimum mating function.

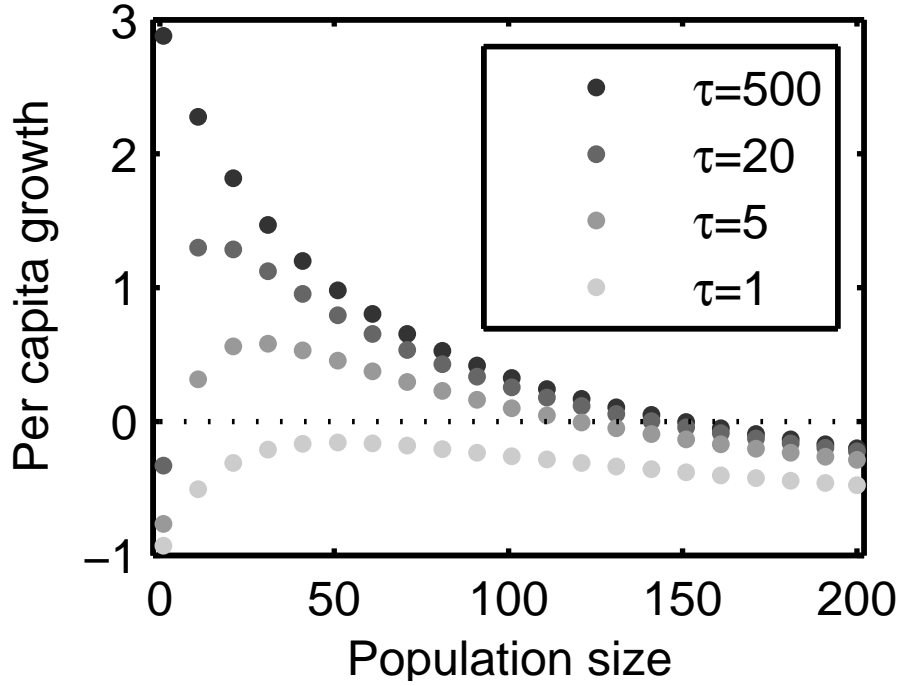
Non-spatial dynamics

Simulations of the non-spatial model (10) show that for short search times (low τ), the unbiased individual-based mating function exhibits a demographic Allee effect, i.e., the per capita growth rate increases with population size, for small populations (Fig. 3). This demographic Allee effect is more pronounced for shorter mate search times (smaller τ), thus τ is effectively an indicator of the Allee effect strength. A population with very short search times ($\tau = 1$) does not grow; the per capita growth is always negative. For very long search times ($\tau = 500$), there is no demographic Allee effect, i.e. the per capita growth decreases monotonically with population size. The other individual-based mating functions (male-biased, female-biased and polygynandrous) all show qualitatively the same patterns as the unbiased individual-based mating function. Similarly, the mass action mating function contains a demographic Allee effect for shorter search times (smaller $c\tau$), and no demographic Allee effect for large search time.

Establishment

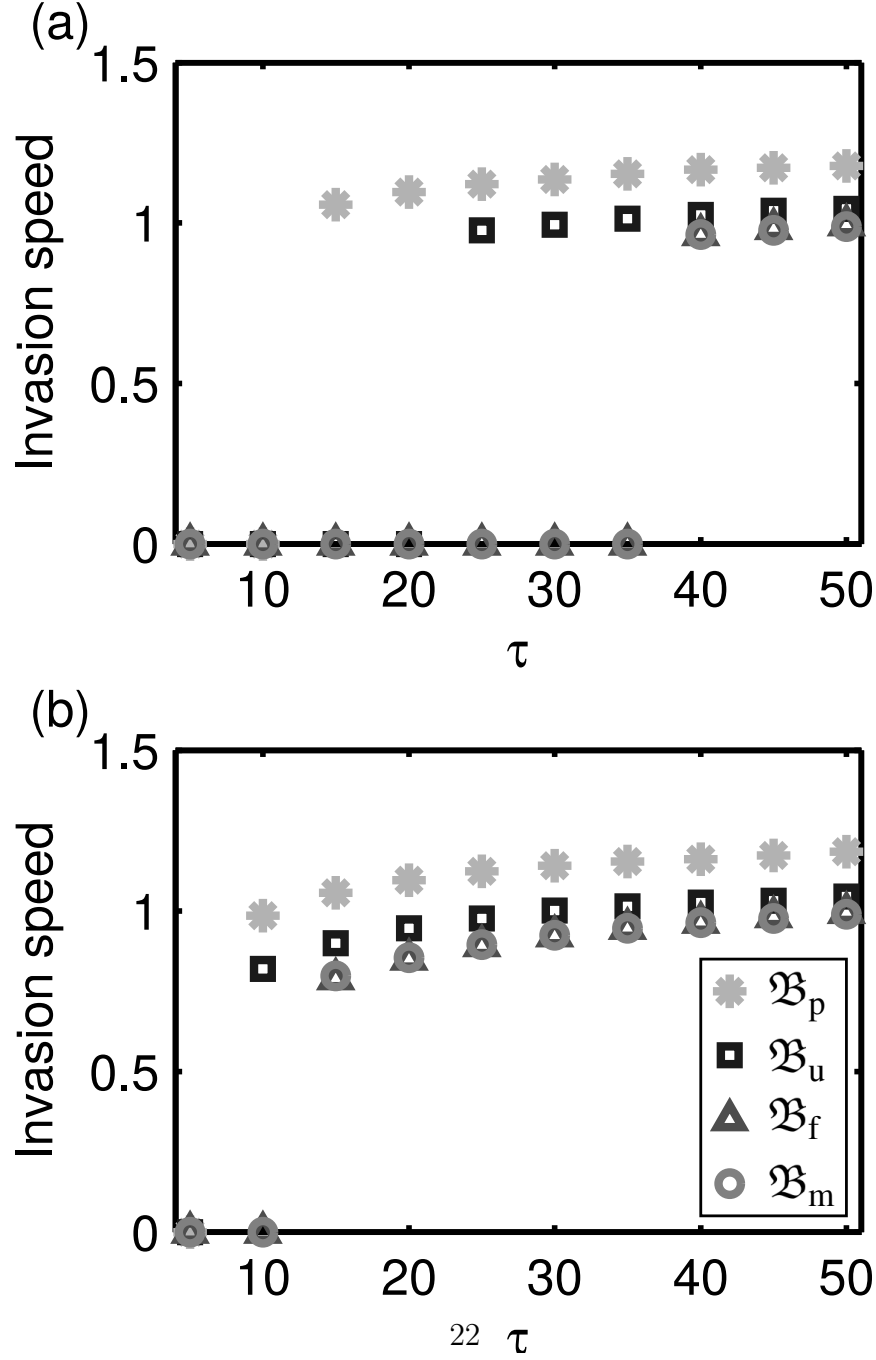
In the spatial spread simulations, populations were only able to establish if they had a sufficiently high initial density, and if individuals could search for sufficiently

Figure 3: The per capita growth rate (eqn. 11) as a function of population size, for simulations of the non-spatial model (eqn. 10) with $b = 100$, $\mu = \phi = 4$, $m = f$ and the individual-based mating function with unbiased searching (\mathfrak{B}_u) and different values of search time, τ .



long mate-search times (τ). The effect of these two conditions was interactive: for larger initial populations, a shorter mate-search time was required for the population to establish (Fig. 4a vs b). This was true for each of the individual-based mating functions (Figs 4, S5). Below a threshold mate-search time and/or initial density, the population failed to establish and went extinct, indicating a strong mate-finding Allee effect. The threshold mate-search time, below which populations went extinct, was higher in spatial simulations than non-spatial simulations. Populations with polygynandrous mating were able to establish under the broadest range of conditions, followed by those with unbiased mating; male-

Figure 4: The long-term invasion speed (averaged across generations 90 to 100) as a function of search time (τ) for (a) initial population radius of $r = 0.15$ and (b) $r = 0.3$, for each of the individual-based mating functions: polygynandrous (\mathfrak{B}_p , stars), unbiased (\mathfrak{B}_u , squares), female-biased (\mathfrak{B}_f , circles), and male-based (\mathfrak{B}_m , triangles). Simulations were run with parameters $b = 1$, $v_m = v_f = 1$ and $\mu = \phi = 4$.

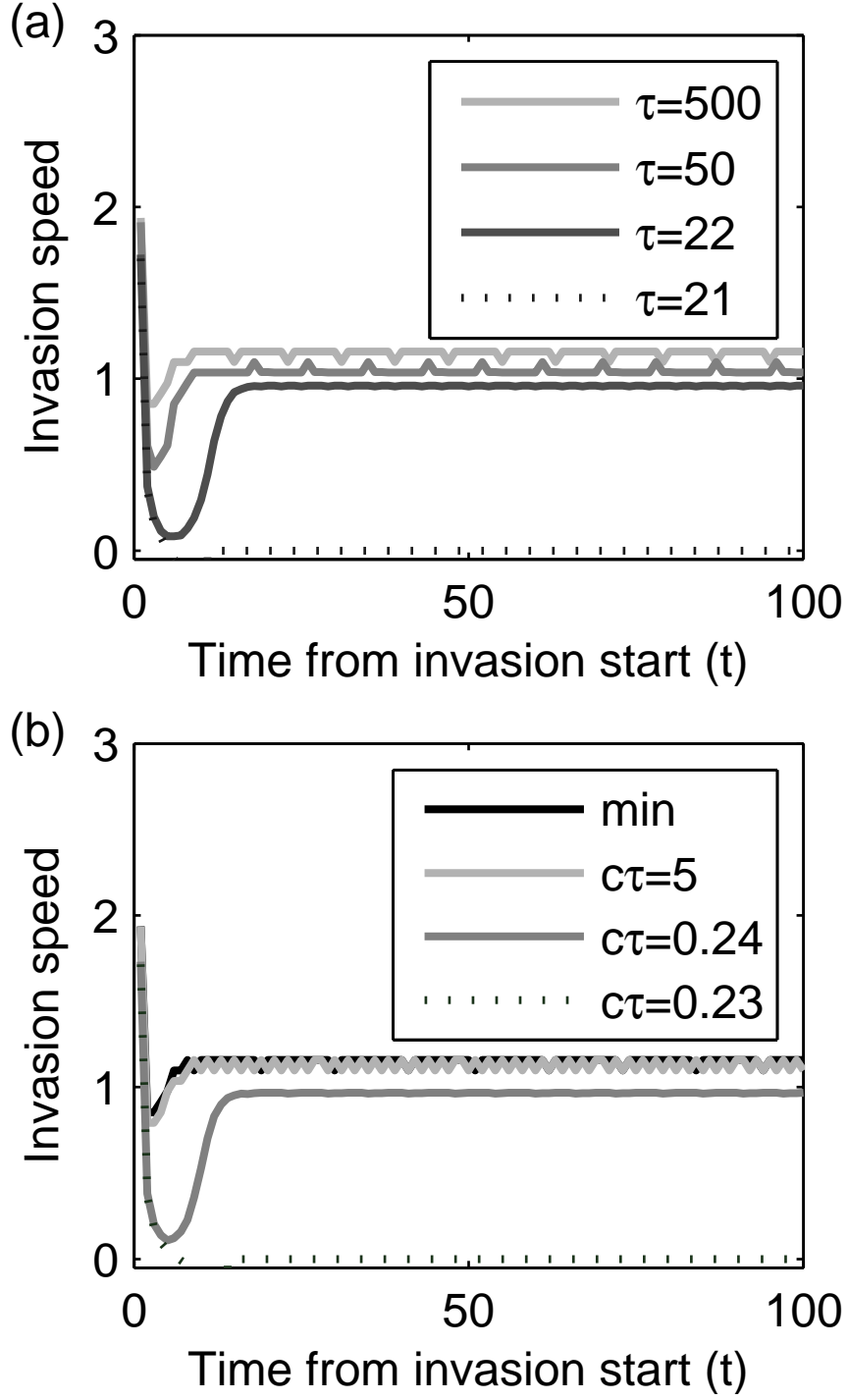


and female-biased mating populations established under the narrowest range of parameter values. Populations with the mass action mating function similarly had a threshold below which they were not able to establish (Fig. S5). The mass action establishment patterns most closely matched those of the unbiased individual mating function. However, populations with the Hölder means mating function were always able to establish and spread, for the range of parameters considered (Fig. S5). Extrapolating the number of matings that occurred at low and high population densities (instead of assuming that the number of matings plateaued) increased the range of conditions where populations could establish (Fig. S6 vs. S5).

Rate of spread

For populations that did successfully establish, the initial population density and mate-search time also affected the spread dynamics. For simulations with the unbiased mating function, the population spread rate started high, dropped to a lower speed, then accelerated to an intermediate rate of spread, which was maintained over time (Fig. 5a). Generally, for larger τ and larger initial densities, both the very initial rate of spread and the lowest spreading speed were higher (Figs S3, S4). For larger τ , the transient acceleration period was shorter and the long-term rate of spread was faster, whereas initial population density had little effect on the long-term spread rate (Fig. S5). Simulations with the male- and female-biased mating functions always spread slightly slower than unbiased mating function simulations, for the same value of τ (Fig. 4). In contrast, simulations with the polygynandrous mating functions always spread faster than those with

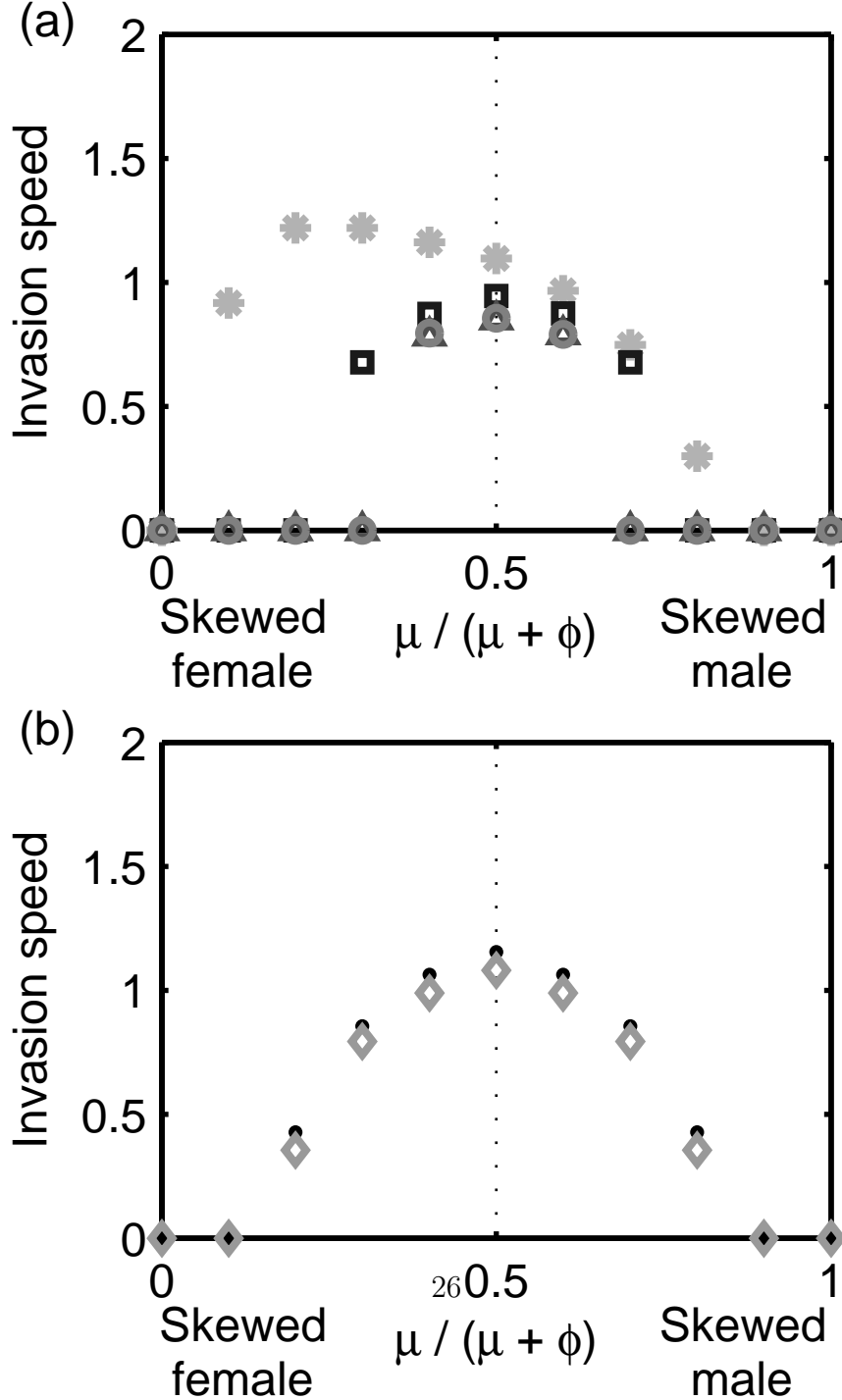
Figure 5: The invasion speed over time for simulations with (a) the unbiased mating function, \mathfrak{B}_u , and (b) the Hölder means (\mathfrak{B}_h) and mass action (\mathfrak{B}_a) mating functions. Parameters: $b = 1$, $r = 0.15$, $v_m = v_f = 1$, $\mu = \phi = 4$.



349 the unbiased mating function, for the same value of τ . A number of simulations
 350 (particularly for large τ) had slight fluctuations in invasion speed, even during
 351 during the long-term rate of spread (Fig. 5). Simulations with the mass action
 352 mating function produced similar results, with populations spreading slowly before
 353 switching to a faster rate of spread (Fig. 5b). For longer mate-search times ($c\tau$),
 354 the period of slow spread was shorter and the long-term rate of spread was faster,
 355 and there were slight fluctuations in the long-term spread rate. Simulations with
 356 the Hölder means mating function with $\alpha = -\infty$ and $\beta = 0.5$ (the minimum mat-
 357 ing function) produced dynamics similar to simulations with the unbiased mating
 358 function and long mate-search times (Fig. 5b). Simulations where the number
 359 of matings was extrapolated at low and high population densities spread slightly
 360 faster than those where the number of matings plateaued, and also did not have
 361 any fluctuations in the long-term rate of spread (Fig. S7 vs 5).

362 Finally, we considered how sex differences in life history parameters influenced
 363 the invasion speed. Here we considered both a skew in sex ratio and sex differences
 364 in dispersal distance. A skewed sex ratio could be the result either of sex ratio
 365 bias at birth, or (more likely) of mortality differences among juveniles. For all
 366 mating functions, as skew in sex ratio at birth increased towards either male or
 367 female (holding the total number of offspring produced per female constant), the
 368 spread rate generally decreased, with a threshold beyond which the invasion failed
 369 (Fig. 6). As with earlier results for the individual-based simulations, populations
 370 with the male-biased and female-biased mating functions were able to establish
 371 under the narrowest range of parameter values and spread the slowest, followed by
 372 unbiased mating functions, with polygynandrous simulations able to spread across
 373 the broadest range of conditions and producing the fastest rates of spread (Fig.

Figure 6: The long-term invasion speed (averaged across generations 90 to 100) as a function of the skew in sex ratio calculated from simulations with the (a) polygynandrous (\mathfrak{B}_p , stars), unbiased (\mathfrak{B}_u , squares), female-biased (\mathfrak{B}_f , circles), and male-based (\mathfrak{B}_m , triangles), individual-based mating functions, and (b) the mass action mating function (\mathfrak{B}_a , diamonds), and the Hölder means mating function (\mathfrak{B}_h , dots). Parameters: $\tau = 20$ (individual-based), $c\tau = 1$ (mass action), $\alpha = -\infty, \beta = 0.5$ (Hölder), $b = 1$, $r = 0.2$, $v_m = v_f = 1$ and $\mu + \phi = 8$.



374 6a).

375 Populations simulated with the mass action and Hölder means mating func-
376 tions produced qualitatively similar results as the unbiased individual-based mat-
377 ing function (Fig. 6b). The polygynandrous mating function was the only one that
378 produced invasion speeds that were asymmetric for male versus female skew in the
379 life history parameters. These simulations spread much faster for female-skewed
380 populations compared to male-skewed ones, presumably because when individuals
381 can mate multiply, sperm (or pollen) is not often limited, whereas the number of
382 mated females becomes the limiting factor for population growth and spread.

383 Skewed sex bias in dispersal (holding the average dispersal distance in the
384 population constant) resulted in qualitatively similar patterns as skewed sex ratio
385 (Fig. S8). Although skewed sex ratio and sex bias in dispersal each alone reduced
386 the rate of spread, combining the two increased the invasion speed in some cases.
387 This occurred only when the skew for each was in opposite directions (e.g., male-
388 biased sex ratio coupled with female-biased dispersal, or female-biased sex ratio
389 coupled with male-biased dispersal) and for very small degrees of bias in each (Fig.
390 S9).

391 Discussion

392 Here we have developed a hybrid approach that links individual mating behaviour
393 to populations dynamics, in a two-sex spreading population. This approach en-
394 ables us to include key aspects of two-sex populations: sex differences in demo-
395 graphic and dispersal parameters, and the mechanics of locating mates. We mod-
396 elled the mate-search process assuming that only males, only females, or both sexes

397 searched for mates, and with monogamous pairs or polygynandrous mating. We
398 determined how these different mating scenarios affected the number of matings,
399 the per capita growth rate, the ability of a population to establish from low den-
400 sity, and the dynamics and rate of population spread. Finally, we compared results
401 from the individual-based mating functions to results from two mating function
402 approximations.

403 Mate-search efficiency, which we here model as the amount of time that indi-
404 viduals had to search for mates (τ) was a key parameter driving population-level
405 outcomes. For low values of τ , there was a strong demographic Allee effect, where
406 the population was not able to grow from low densities. For slightly larger values
407 of τ , a non-spatial population that grew without spreading was able to establish,
408 while adding spatial spread caused establishment to fail. (Note that this is a typi-
409 cal result for spatial spread with an Allee effect; [Lewis & Kareiva 1993](#); [Goodsman
410 & Lewis 2016](#).) For larger values of τ , the demographic Allee effect became weaker,
411 and the population was able to establish in both spatial and non-spatial contexts,
412 but went through a long transient phase of slow spread before switching to a faster
413 long-term spread rate. For even larger values of τ , the demographic Allee effect
414 vanished, the slow phase of population spread was quite short, and the long-term
415 rate of spread was quite fast. These results held within each of the individual-based
416 mating functions considered: male- or female-biased mate searching with monog-
417 amous pair formation, and unbiased mate searching with monogamous pairs or
418 polygynandrous mating. However, populations with sex-biased mate searching
419 were the least robust, failing to invade for conditions under which unbiased mate
420 searching populations were successfully able to invade, and spreading at a slower
421 rate when they were able to invade. In contrast, populations with polygynandrous

422 mating were the most robust, able to spread in conditions where populations with
423 each of the monogamous mating functions were not able to establish. Initial popu-
424 lation density had some interactive effects with mate search time: populations that
425 started at too low a density failed to establish, and as initial population density
426 increased the threshold mate search time required for establishment decreased.

427 The dynamics were to some extent found to be different between the individual-
428 based mating functions and the approximations. This is unsurprising per se, as
429 approximations do not perfectly capture the mating dynamics in the IBM, but it
430 is worthwhile to comment on when precisely large failures were found. The mass
431 action mating function was a good proxy for the case of monogamous mating if
432 searching had no sex bias, but it performed less well if searching was sex-biased or
433 if matings were polygynandrous. The Hölder means mating function was only a
434 good proxy for the monogamous mating function for very large search times (big
435 τ). Even though the effect of the Hölder means' α parameter on the number of mat-
436 ings was qualitatively similar to increasing search time τ in the individual-based
437 simulations, the Hölder means mating function did not display any of the Allee
438 threshold behaviour that was characteristic of the dynamics of the individual-based
439 mating functions. The Hölder means mating function also shared the property of
440 the mass action function that it did not form a good proxy when searching was
441 sex-biased or there was polygynandry. For these reasons, some dynamic effects
442 were only discoverable using the IBM.

443 Some of our simulations showed fluctuations in the long-term invasion speed,
444 beyond the initial period of slow spread. The fluctuations were only noticeable if
445 we assumed that the number of matings plateaued at low and high densities (rather
446 than linearly extrapolating the number of matings at low and high densities). At

low densities, this is the equivalent to imposing a strong mate-finding Allee effect (a threshold density below which no matings, and therefore no population growth, occurs). However, a strong mate-finding Allee effect alone is not sufficient to generate these oscillations. For comparison, we ran simulations with a minimum mating function that was calculated exactly at all densities (no interpolation) but with a threshold density below which no matings occurred. These simulations produced smaller oscillations (Fig. S10), suggesting that oscillations are caused by some combination of a strong mate-finding Allee effect and the nonlinearities introduced by interpolating the mating function at intermediate densities. Recent work on oscillations in invasion speed in one-sex models supports this idea (Sullivan et al. unpublished).

Our findings broadly align with past empirical and theoretical work demonstrating that Allee effects can prevent (Lewis & Van Den Driessche, 1993; Goodman & Lewis, 2016), slow (Lewis & Kareiva, 1993; Taylor & Hastings, 2005; Tobin et al., 2007; Krkošek et al., 2012) or delay (Hopper & Roush, 1993; Kot, Lewis & van den Driessche, 1996; Veit & Lewis, 1996) population spread. The novelty of our study is showing how the details of individual behaviour determines the impact of a mate-finding Allee effect. Our work also builds on the models of Miller et al. (2011) and Shaw & Kokko (2015), which separately examined the influence of sex-biased dispersal and of mating dynamics, respectively. Here we have shown that sex-specific life history traits (like sex-biased dispersal) interact with mating mechanics, typically resulting in both a reduced invasion probability and a slower invasion speed. The exception is when individuals can mate multiple times and when life history traits are skewed female (females disperse farther or female-skewed sex ratio): with this combination the population spreads faster than

a population with only one of these factors. We have also shown that different sex-specific life history traits have compensatory effects on invasion speed: skewed sex ratio and sex-biased dispersal alone each slow an invasion, but a male-skewed sex ratio combined with female-biased dispersal (and vice versa) can speed up an invasion.

Our model describes mate-finding difficulties by altering the time, τ , available for individuals to search before they have to give up and remain unmated (for life, as ours is a discrete generation model). The longer the time is, the wider the area that the individuals can search, thus τ in our model can be interpreted in several possible ways, broadening the scope of the model from the narrowest interpretation that τ must be measured in units of time. Under a more general view, high τ reflects mate-search efficiency that can be brought about by efficient locomotion and/or high sensory capacities to locate mates that are relatively far away. For example, male *Lymantria dispar* (gypsy moths) find females that emit pheromones; interestingly, despite mate searching therefore presumably being relatively efficient, it still is not efficient enough to avoid Allee effects at range boundaries (Contarini *et al.*, 2009). Fish lice (parasites of fish) such as *Lepeophtheirus salmonis* also appear to be intermediate cases: males disperse among fish in search of mates, but again, this is not efficient enough to avoid mate-finding Allee effects (Krkošek *et al.*, 2012).

When do we expect τ to be low enough in natural systems to cause mate-finding Allee effects? Our model is not evolutionary (τ does not evolve), but evolutionary ideas can help elucidate when invasion speeds are likely to be negatively impacted by mate availability. If females in the core of their range do not remain unmated, τ is ‘large enough’ (i.e., leads to high female fertilization rate) under these conditions.

497 Newly established populations may then ‘use’ the same value for τ even if it now
498 leads to poorer mating success — either because there’s not sufficient heritable
499 variation present to lead to improved behaviours, or because of gene flow from the
500 core range importing behaviours that create a too low τ for the local conditions.
501 We should note, however, that not all marginal populations necessarily experience
502 mate-finding difficulties (Fauvergue & Hopper, 2009); a possible cause is that
503 range margins might maintain high local densities despite being more fragmented
504 (Gaston 2009, however note that this paper is not focused on an explicit invasion
505 context).

506 Mate-finding Allee effects are typically defined as a reduced ability, or failure,
507 to find mates at low density (Dennis, 1989; Wells *et al.*, 1998; Stephens, Sutherland
508 & Freckleton, 1999). Although there is consensus on this general definition, ex-
509 plicit criteria for determining whether a system displays a mate-finding Allee effect
510 are rarely given. Courchamp, Berec & Gascoigne (2008) provide one such set of
511 criteria: the probability of mating per female should increase (or be constant) as
512 the number of males increases, and should decrease (or be constant) as the num-
513 ber of females increases. However, within this definition, the effect of an overall
514 increase in density (with a constant ratio of males to females) on probability of
515 mating per female is not specified. To address this deficiency, and to align the
516 explicit criteria for mate-finding Allee effects with the general definition, here we
517 propose that an additional criterion be included. Namely, we suggest that at low
518 density, increasing the population density (while keeping the sex ratio constant)
519 should increase a female’s probability of mating. However, a number of com-
520 monly used mating functions violate this criterion, meaning they do not display
521 a mate-finding Allee effect under our revised definition. Specifically, within the

demography literature, a mathematically desirable (albeit controversial; Iannelli, Martcheva & Milner 2005) property of any mating function is that it should be homogeneous of degree one (Yellin & Samuelson, 1974; Legendre, 2004), that is, changing the population size (with a constant ratio of males to females) changes the number of matings, proportionally. This is generally equivalent to stating that a female’s probably of mating is independent of population size, which conflicts with the above general definition of a mate-finding Allee effect.

Future studies could expand on the methods we develop here in order to explore how details of individual behaviour influence population spread rates. First, a wider range of individual mate finding behaviour could be explored, such as using a deterministic (instead of probabilistic) stopping rule or using signalling or homing to locate mates (Vane-Wright & Boppre, 1993). Second, many organisms have other adaptations for alleviating mate-finding Allee effects, such as moving more frequently at low density (Kindvall *et al.*, 1998), storing sperm across mating seasons (Parker, 1970), and context-dependent sex determination (Becheikh *et al.*, 1998). Finally a wider range of interactions between mating and life history characteristics could be considered – e.g., although we explored the effects of polygynandrous mating and sex-biased searching, we did not explore the effect of both factors combined. Each of these mechanisms could be simulated to determine how they influence population growth and spread rate.

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550 Authors' Contributions

551 AKS and MGN conceived the model idea. AKS developed the model, ran simula-
552 tions and wrote the paper, with contributions from HK and MGN.

553 Data accessibility

554 Model code is available from the Dryad Digital Repository: <http://doi.org/10.5061/dryad.39n4g>
555 ([Shaw, Kokko & Neubert, 2016](#)).

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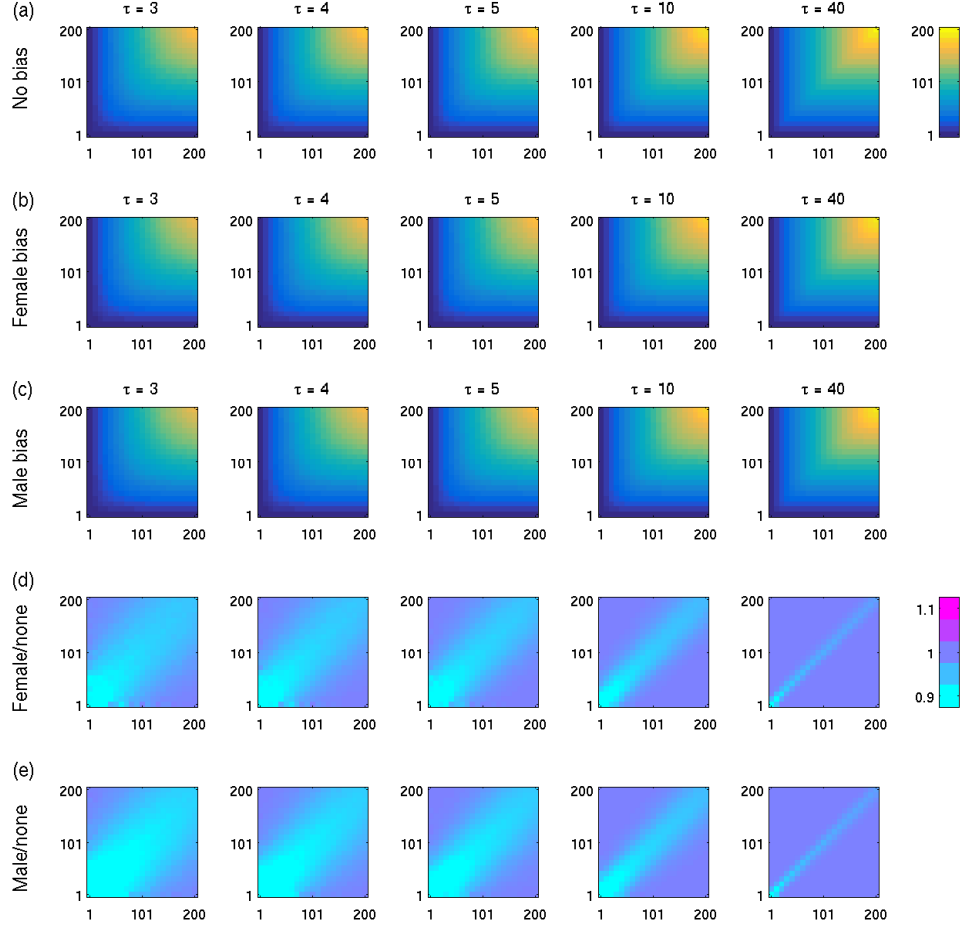


Figure 7: The number of pairs formed as a function of the number of females (x-axis) and males (y-axis) present, comparing across three different individual-based simulations for different values of τ (the number of time steps allowed for pair formation). (a) The number of pairs formed when both males and females search (\mathfrak{B}_u). (b) The number of pairs formed when only females search and males are stationary (\mathfrak{B}_f). (c) The number of pairs formed when only males search and females are stationary (\mathfrak{B}_m). (d) The ratio between the values in (b) and (a). (e) The ratio between the values in (c) and (a). All mating functions were calculated using the mate search shape parameter $a = 5$.

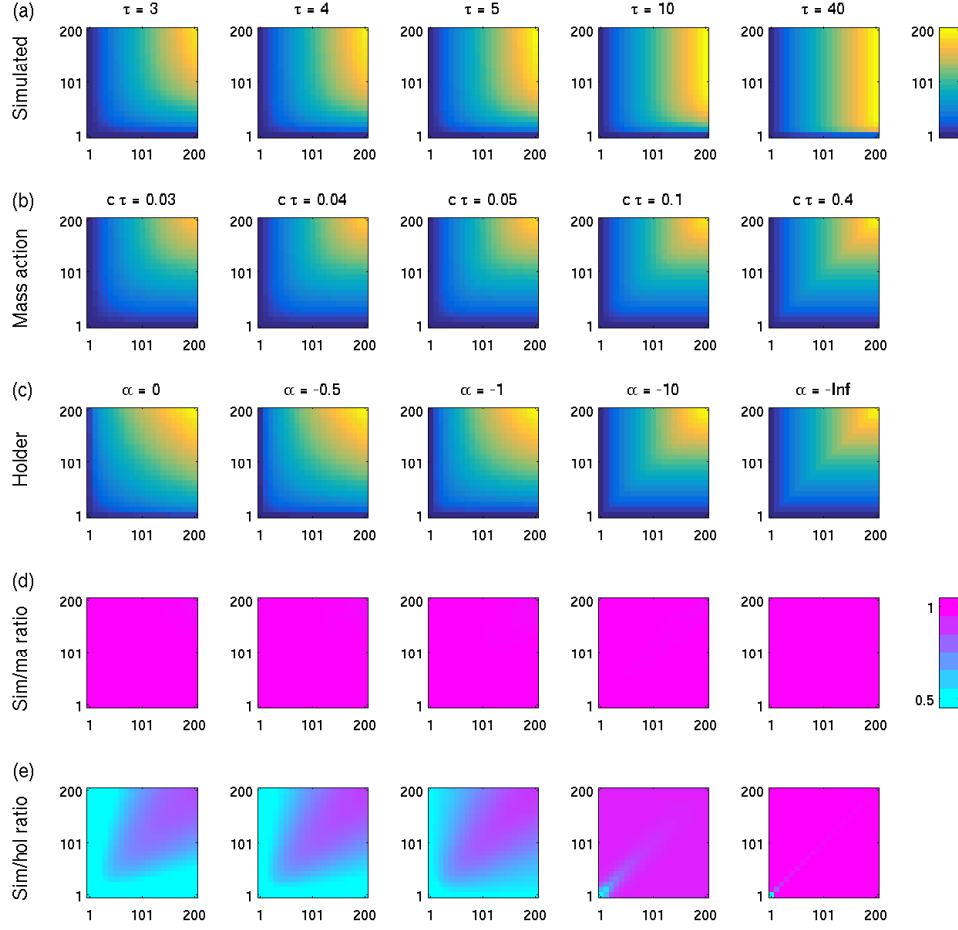


Figure 8: The number of pairs formed as a function of the number of females (x-axis) and males (y-axis) present, comparing across different mating functions. (a) The number of pairs formed in the individual-based mating function with polygynandrous mating (\mathfrak{B}_p), for different values of τ (the number of time steps allowed for pair formation). (b) The number of pairs formed in the mass action mating function (\mathfrak{B}_a) for different values of $c\tau$. (c) The number of pairs formed in the Hölder means mating function (\mathfrak{B}_h) with $\beta = 0.5$ and different values of the shape parameter α . (d) The ratio between the values in Fig. S1a and (b) here. (e) The ratio between the values in Fig. S1a and (c) here. All mating functions were calculated using the mate search shape parameter $a = 5$.

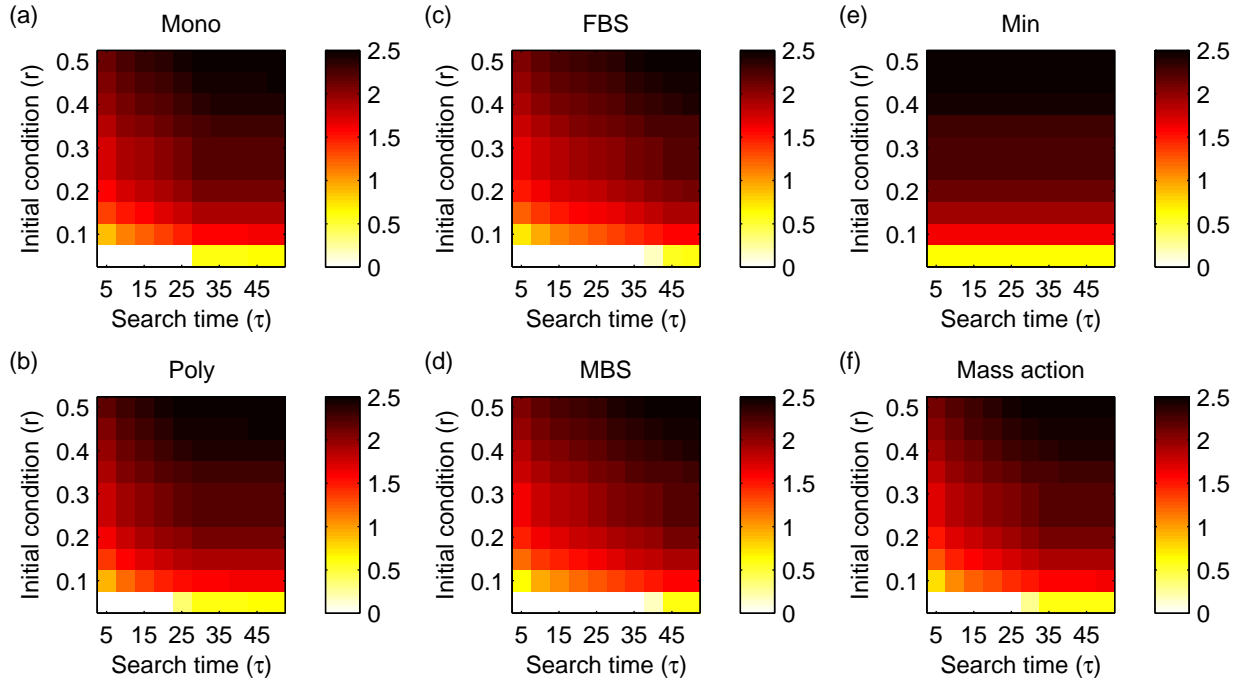


Figure 9: The initial invasion speed (generation 1) as a function of search time (τ) and the simulation initial population radius (r), for each of the mating functions. White regions indicate a failed invasion. Simulations were run with parameters $b = 1$, $v_m = v_f = 1$ and $\mu = \phi = 4$. The number of matings were assumed to plateau in the mating functions.

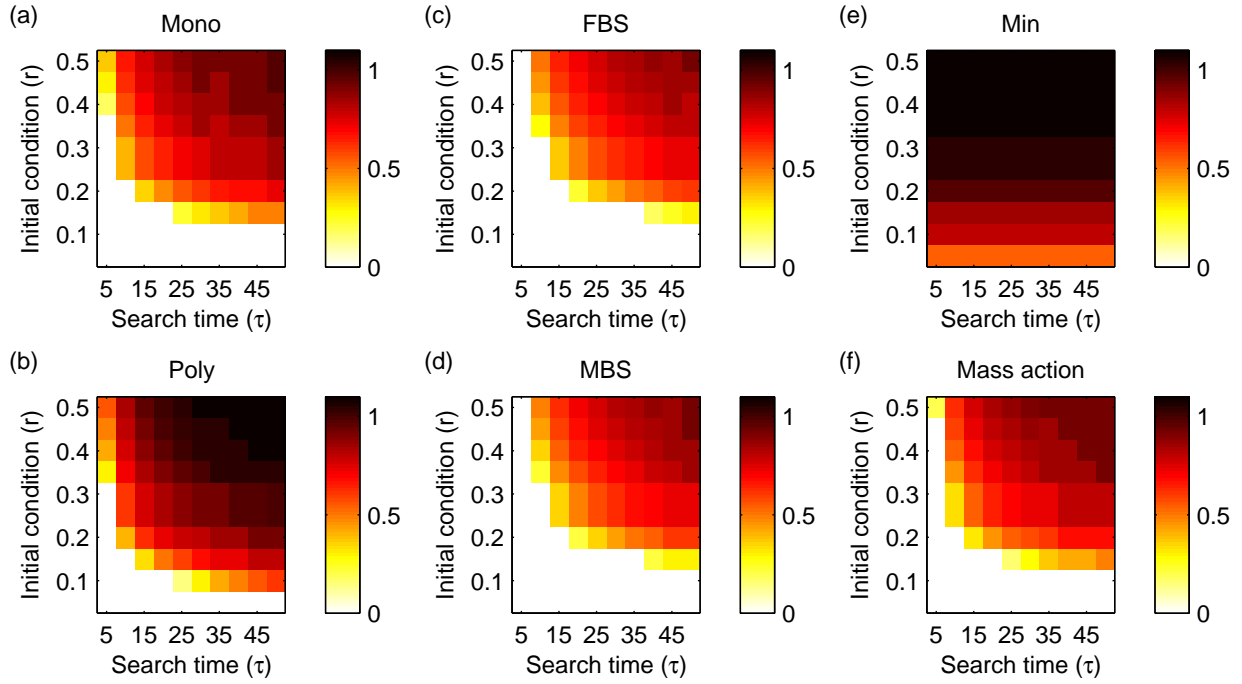


Figure 10: The minimum invasion speed as a function of search time (τ) and the simulation initial population radius (r), for each of the mating functions. White regions indicate a failed invasion. Simulations were run with parameters $b = 1$, $v_m = v_f = 1$ and $\mu = \phi = 4$. The number of matings were assumed to plateau in the mating functions.

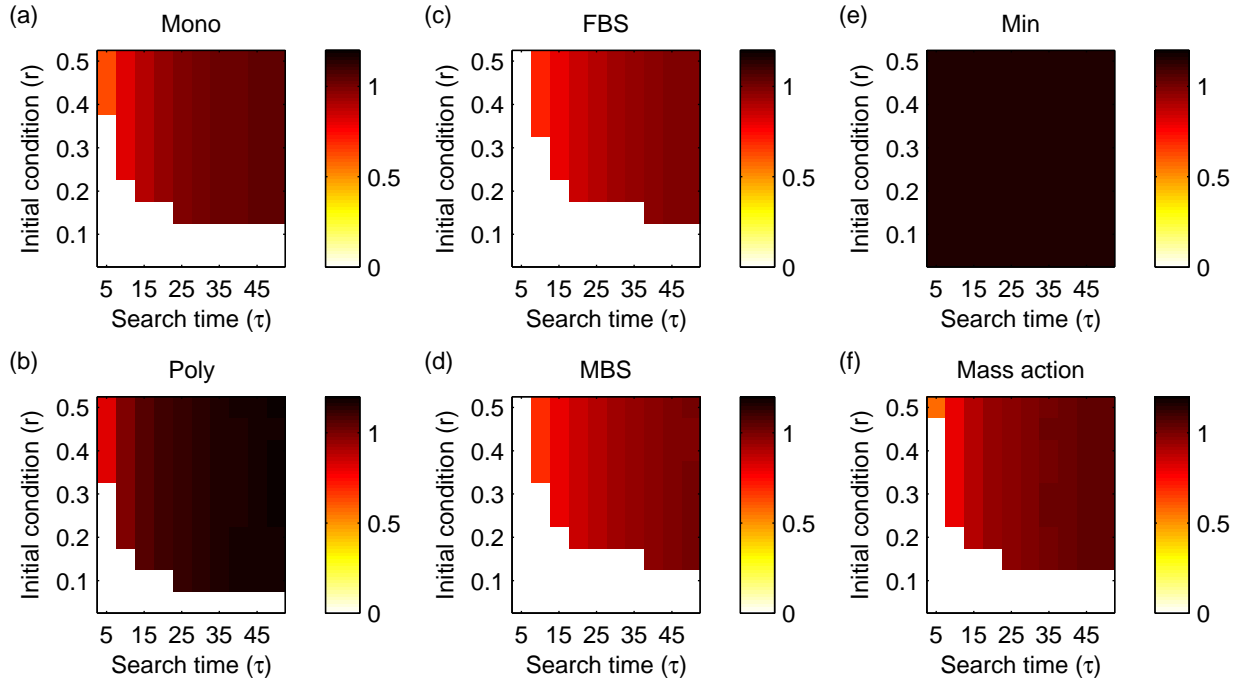


Figure 11: The long-term invasion speed (averaged across generations 90 to 100) as a function of search time (τ) and the simulation initial population radius (r), for each of the mating functions. White regions indicate a failed invasion. Simulations were run with parameters $b = 1$, $v_m = v_f = 1$ and $\mu = \phi = 4$. The number of matings were assumed to plateau in the mating functions.

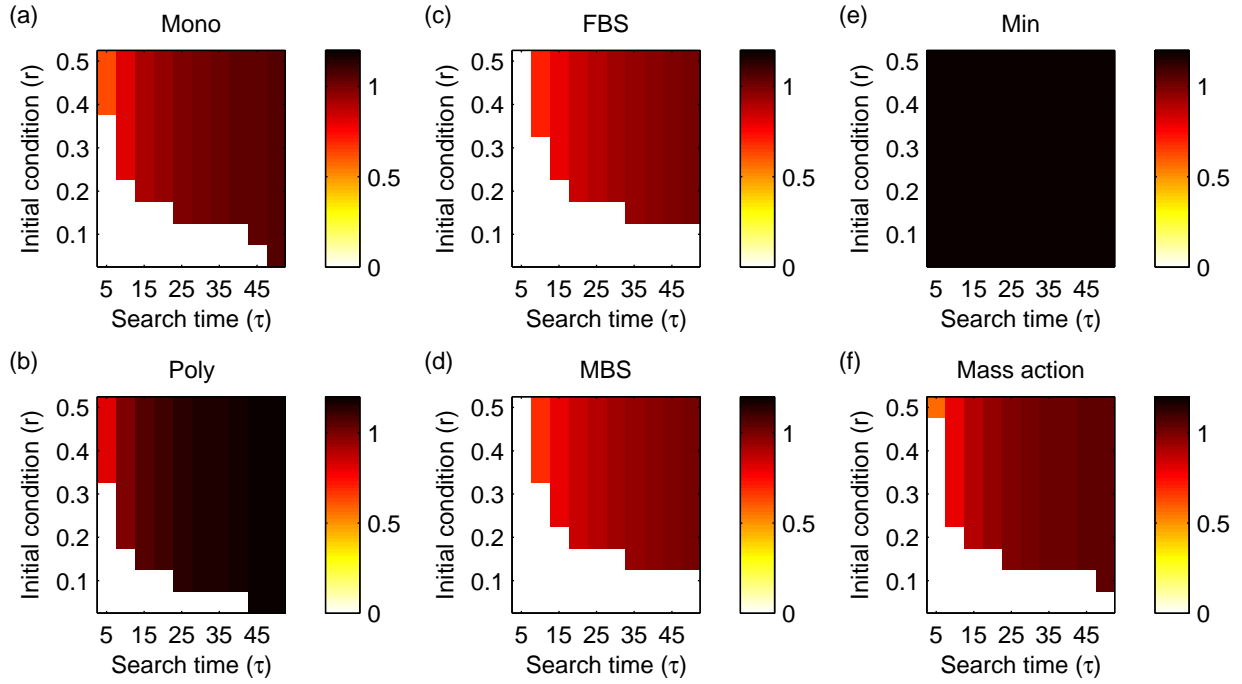


Figure 12: The long-term invasion speed (averaged across generations 90 to 100) as a function of search time (τ) and the simulation initial population radius (r), for each of the mating functions. White regions indicate a failed invasion. Simulations were run with parameters $b = 1$, $v_m = v_f = 1$ and $\mu = \phi = 4$. The number of matings were extrapolated at high and low densities in the mating functions.

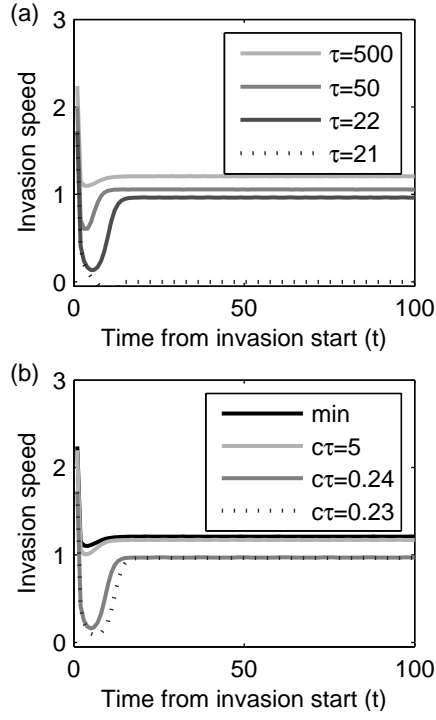


Figure 13: The invasion speed over time for simulations with (a) the unbiased mating function, \mathfrak{B}_u , and (b) the Hölder means (\mathfrak{B}_h) and mass action (\mathfrak{B}_a) mating functions. Parameters: $b = 1$, $r = 0.2$, $v_m = v_f = 1$, $\mu = \phi = 4$. The number of matings were extrapolated at high and low densities in the mating functions.

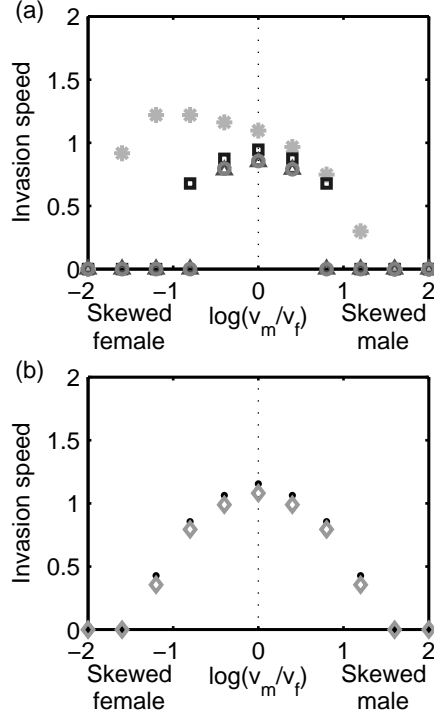


Figure 14: The long-term invasion speed (averaged across generations 90 to 100) as a function of the degree of sex bias in dispersal, calculated from simulations with the (a) polygynandrous (\mathfrak{B}_p , stars), unbiased (\mathfrak{B}_u , squares), female-biased (\mathfrak{B}_f , circles), and male-biased (\mathfrak{B}_m , triangles), individual-based mating functions, and (b) the mass action mating function (\mathfrak{B}_a , diamonds), and the Hölder means mating function (\mathfrak{B}_h , dots). Parameters: $\tau = 20$ (individual-based); $c\tau = 1$ (mass action); $\alpha = -\infty, \beta = 0.5$ (Hölder); with $b = 1, r = 0.2, v_m + v_f = 2$ and $\mu = \phi = 4$.

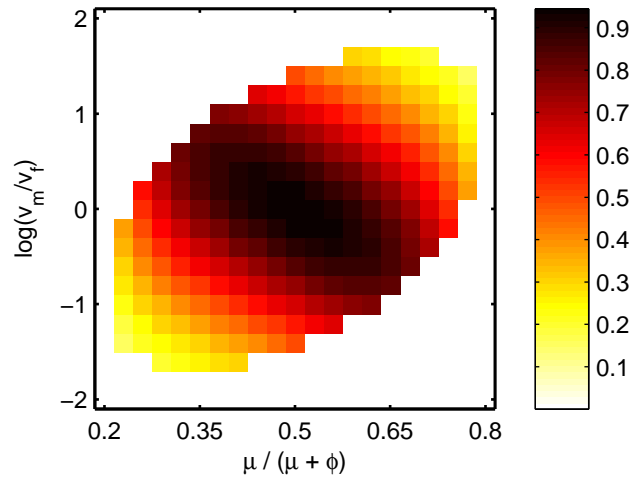


Figure 15: The long-term invasion speed (averaged across generations 90 to 100) as a function of the skew in sex ratio (x-axis) and the degree of sex bias in dispersal (y-axis), calculated from simulations with the unbiased individual-based mating function, \mathfrak{B}_u . Parameters: $\tau = 20$, $b = 1$, $r = 0.2$, $v_m + v_f = 2$ and $\mu + \phi = 8$.

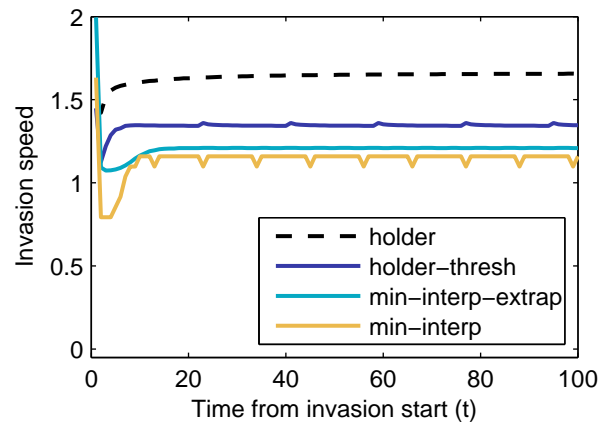


Figure 16: The invasion speed over time for simulations with different levels of interpolation/extrapolation.